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## INSTRUMENTAL VARIABLES IN ACTION

Two Examples

#### Swanz Ganz Cathether

- Swan-Ganz catheter, which is a device used by intensive care unit (ICU) doctors to guide therapy
- A Swan-Ganz catheter is a long slender tube outfitted with sensors designed to measure hemodynamic pressures in the right side of the heart and in the pulmonary artery.



#### Swan Ganz Catheter

Under sterile conditions, an ICU doctor will typically insert the catheter into the left subclavian vein (underneath the clavicle)







#### Swan-Ganz catheter

- A Swan-Ganz catheter is primarily used to measure pressure at different locations in the right side of the heart.
- These pressure measurements can provide important diagnostic information, such as whether the patient's heart valves are working, or whether the patient has pulmonary hypertension.
- Information gleaned from Swan-Ganz measurements is often used by ICU doctors to make decisions about treatment, such as whether to give the patient medications that affect the functioning of the heart.



- The procedure is done to evaluate how the blood moves (circulates) in people who have:
- □ Abnormal pressures in the heart arteries
- □ <u>Burns</u>
- □ <u>Congenital heart disease</u>
- □ <u>Heart failure</u>
- Kidney disease
- Leaky heart valves (valvular regurgitation)
- □ <u>Shock</u>



- Swan-Ganz catheterization can also be used to detect abnormal blood flow between two areas of the heart that are not normally connected.
- Conditions that can also be diagnosed or evaluated with Swan-Ganz catheterization include:
- □ <u>Cardiac tamponade</u>
- Pulmonary hypertension
- Restrictive cardiomyopathy



It may also be done to monitor for complications of <u>heart attack</u> and to see how well certain heart medicines are working.



#### **Swan-Ganz Catheter & Econometrics**

- Connors et al. (1996), which examines the impact of Swan-Ganz catheterization on mortality outcomes among a population of patients admitted to the intensive care unit (ICU) at five prominent hospitals.
- Connors et al. (1996) reach the controversial conclusion that patients who receive Swan-Ganz catheterization during their first day in the ICU are 1.27 times more likely to die within 180 days of their admission.



- Before Connors et al. (1996), Gore et al. (1985) and Zion et al. (1990) also found found that catheterization increases mortality. Dalen (2001) criticized both studies because they did not control for clinically important differences between the patients who had catheters placed and those who did not.
- The Connors et al. (1996) study was conceived in part as a response to this criticism.



#### Dataset

The "outcome" variables are:







#### The main regressor is

• swang = 
$$\begin{cases} 1 & \text{if patient received SG catheter} \\ 0 & otherwise \end{cases}$$



### A first "stunning" regression

□ We run the following regressions:

$$surv7_{i} = \beta_{0} + \beta_{1}swang_{i} + u_{i}$$
$$surv30_{i} = \beta_{0} + \beta_{1}swang_{i} + u_{i}$$

$$surv180_i = \beta_0 + \beta_1 swang_i + u_i$$

#### • What is the interpretation of $\beta_1$ in each model?



#### **Estimated Effects of Catheterization**



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#### Results

- Our simple analysis seems to confirm the results of the first generation studies
- Catheterization increases the probability of death by 5%-8%, depending on the time span.
- Catheterization also increases the probability of death at 7 days from admission...



There are however potential problems





#### □ There is big concern:

• Omitted variables.....



#### **Omitted Variable**

- Catheterization may depend on other personal characteristic that affect both catheterization decision and outcome
- Sicker patients are probably catheterized more frequently than relatively healtier patients



## Adding Regressors

We have a wealth of information about the patients:

- Personal characteristic
  - age, sex, education, race, education, income, etc.
- □ Categories of admission diagnosis
  - respiratory, cardiovascular, neurological, gastrointestinal, renal, metabolic, hematologic, etc.
- Categories of comorbidities illness:
  - hearth failure, depression, cirrhosis, tumor, HIV, etc.
- Other information
  - estimate of the prob. of surviving 2 months



#### **New Model**

#### □ We estimate:

$$surv7_{i} = \beta_{0} + \beta_{1}swang_{i} + x_{i}\gamma + u_{i}$$
$$surv30_{i} = \beta_{0} + \beta_{1}swang_{i} + x_{i}\gamma + u_{i}$$

 $surv180_i = \beta_0 + \beta_1 swang_i + x_i\gamma + u_i$ 

:



## **Additional Problem**

- Recall that the outcome variables are binary
   surv7 is either 1 (patient is alive after 7 days) or 0
- The linear probability model could not be appropriate
  - We should use Probit and Logit to verify the robustness of the results
  - In this case we have to be careful in interpreting the results...



#### "Full" Linear Probability Model

$$surv7 = 0.6306 - 0.0150 \times swang + \dots$$
  

$$surv30 = 0.2709 - 0.0587 \times swang + \dots$$
  

$$surv60 = 0.1572 - 0.0648 \times swang + \dots$$
  

$$(0.0825) - (0.0131) \times swang + \dots$$
  

$$surv90 = 0.1309 - 0.0625 \times swang + \dots$$
  

$$(0.0827) - (0.0609 \times swang + \dots)$$
  

$$surv120 = 0.0970 - 0.0609 \times swang + \dots$$
  

$$(0.0830) - (0.0133) \times swang + \dots$$
  

$$surv150 = 0.0728 - 0.0617 \times swang + \dots$$
  

$$(0.0829) - (0.0133) \times swang + \dots$$
  

$$(0.0829) - (0.0575 \times swang + \dots)$$
  

$$(0.0829) - (0.0134) \times swang + \dots$$



## **Comparison of the results**

#### No controls



#### With controls



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- □ When we add controls the effects tend to become smaller in magnitude, but they are still negative
- Except for surv7, they are still statistically different from 0
- The effect of catheterization ranges from 5 to 6 percentage points



# The unconditional probability of dying within 180 days from admission in ICU is 50.67%

## With a Swang-Ganz catheter 56.42%

#### □ 56.42/ 50.67 = 1.11 more likely to die



## **Probit and Logit**

□ We estimate now the following model:

$$Pr(surv7_i = 1 | swang_i, x_i) = \Phi(\beta_0 + \beta_1 swang_i + x'_i\gamma)$$
$$Pr(surv30_i = 1 | swang_i, x_i) = \Phi(\beta_0 + \beta_1 swang_i + x'_i\gamma)$$

 $\Pr(surv180_i = 1 | swang_i, x_i) = \Phi(\beta_0 + \beta_1 swang_i + x'_i \gamma)$ 

□ The interpretation of the coefficient is different..



## **Probit/Logit Estimated Coefficients**

#### **Probit**

#### Logit





## **Probit/Logit Effects**

**Probit - Effects** 

**Logit - Effects** 

□ Estimate

□ Estimate

$$\Phi(\hat{\beta}_0 + \hat{\beta}_1 + \bar{x}'\hat{\gamma}) - \Phi(\hat{\beta}_0 + \bar{x}'\hat{\gamma}) = F(\hat{\beta}_0 + \hat{\beta}_1 + \bar{x}'\hat{\gamma}) - F(\hat{\beta}_0 + \bar{x}'\hat{\gamma})$$

where

$$\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$$



## **Probit/Logit Estimated Effects**

**Probit - Effects** 

#### **Logit - Effects**





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#### Using Probit and Logit instead of the Linear

□ Probability model does not change the message:

- Catheterization has a negative effect on the probability of surviving
- The Probit effects are very close to the LPM effects
- The Logit effects are somewhat smaller, yet still significantly different from 0



### **Unobserved Omitted Variable**

- A threat to the internal validity of the study is the possibility that doctors do have information on health status of the patient
- This information may not be contained in the dataset
- If it is so, there is an 'unobserved' omitted variable that might render the interpretation of the effects as casual invalid



We could potentially solve the "endogeneity" caused by the unobserved omitted variable using instrumental variable technique.

□ How do we find a valid instrument for this study?

□ Think, think....



#### Two conditions for a valid instrument

## For an instrumental variable (an "*instrument*") *Z* to be **valid**, it must satisfy two conditions:

Instrument relevance
Instrument exogeneity



## Instrumental Variable

An instrument is a random variable Z
 E[Z'u]=0 (exogenous)
 It enters the first stage regression (relevance)
 surv7<sub>i</sub> = β<sub>0</sub> + β<sub>1</sub>swang<sub>i</sub> + x'<sub>i</sub>γ + u<sub>i</sub>

endogenous



exogenouss





There are more catheterization on weekdays:
 ✓ Tuesday, Wednesday, Thuersday, Friday

#### □ We consider the following instrument

$$\square Z_i = \begin{cases} 1 & \text{if day of admission is Su, Sa, Mo} \\ 0 & otherwise \end{cases}$$

- □ Two questions:
  - Is it exogenous?
    Is it relevant?



## Exogeneity of day of the week

- Exogeneity means that the day of the week is uncorrelated with the unobserved health
- There is not a priori reason to believe that the day of the week is correlated with unobserved health conditions
- We cannot test this assumption since the model is exactly identified.



#### Relevance

The instrument must be correlated (in the first stage sense) with swang

□ The idea is that

On weekends there are fewer ICU doctors on the premises and the probability of being catheterized is much lower;

There is evidence that staffing affects treatment decisions



#### **Testing relevance**

#### **First Stage**

#### □ Simple regression:

 $swang = \underset{(0.0076)}{0.0264} \times Z$ 

#### □ First Stage

 $swang = 0.6175 - 0.0271 \times Z + x'_{i}\hat{\delta}$ 

□ t-test:

$$\frac{0.0271}{0.0129} = 2.10 > 1.96$$

Statistically  $\neq 0$  at 5%



Testing

#### TSLS

#### □ Predict *swang*

$$\widehat{swang}_i = \underset{(0.0824)}{0.0129} - \underset{(0.0129)}{0.0271} \times Z_i + x_i' \hat{\delta}$$

#### Second Stage(s)

$$surv7_{i} = \beta_{0} + \beta_{1}\widehat{swang}_{i} + x_{i}\gamma + u_{i}$$
$$surv30_{i} = \beta_{0} + \beta_{1}\widehat{swang}_{i} + x_{i}\gamma + u_{i}$$

$$surv180_i = \beta_0 + \beta_1 \widehat{swang}_i + x_i \gamma + u_i$$



:

#### **TSLS Results**





## Comments

#### □ TSLS:

 The effects are estimated to be much (implausibly) larger

 Catheterization seems to reduce mortality at shorter time horizon (7 days)

These effects are all statistically insignificant (not different from 0)



#### Caveats

## The first stage involves a binary variables The predictions

$$\widehat{swang}_i = \underset{(0.0824)}{0.0129} - \underset{(0.0129)}{0.0271} \times Z_i + x_i' \hat{\delta}$$

Can be outside the [0,1]

- There are way to fix this problem by using a "probit instrumental variable" method.
- □ Weak instruments.....



#### Weak instruments

□ The rule-of-thumb says that we should worry if the *Wald-stat* >10m

What is the Wald-stat in our first stage?
Hint 1: there is only one endogenous regressor
Hint 2: for testing only one coefficient the Wald-test is numerically equivalent to ....



#### Effect of Education on Wages

- The objective is to estimate the "causal" effect of education on individual wages
- Very difficult to do with OLS because of omitted ability
- This is usually addressed by Instrumental Variables techniques



 D. Card (1995), "Using Geographic Variation in College Proximity to Estimate the Return to Schooling," in Aspects of Labour Market Behavior: Essays in Honour of John Vanderkamp. Ed. L.N. Christophides, E.K. Grant, and R. Swidinsky, 201-222. Toronto: University of Toronto Press.



#### Data

## Sample of 3009 workers from National Longitudinal Survey

http://www.bls.gov/nls/

□ Variables:

 Hourly wage (lwage), age (age), race (black), experience (exper), geographic information (smsa), family background (fatheduc, motheduc), standardized test scores (IQ, KWWW).



#### Data

#### MSA – Metropolitan Statistical Area

- In the United States, a metropolitan area refers to a geographical region with a relatively high population density at its core and close economic ties throughout the area.
- □ KWW Knowledge of the World of Work

A direct measure of "ability"

□ IQ

Measure of intelligence.



#### **Omitted Variable Bias**

#### $log(wage) = \beta_0 + \beta_1 educ + u$

#### "Ability Bias"

Suppose that some individuals have an an unobserved characteristic ("ability") that enables them to earn higher wages at any level of education. If these individuals acquire higher than average schooling then the OLS estimate of β<sub>1</sub> will be upward-biased



#### **Basic regression**

> summary rob(lm(lwage~educ, data=cd))

Dependent varaible: lwage

Coefficients:

	Estimate	Std.	Error	t	value	Pr(> t )
(Intercept)	5.57088	0.	03914		142.4	<2e-16
educ	0.05209	0.	00291		17.9	<2e-16

\_\_\_

Heteroskadasticity robust standard errors used

Residual standard error: 0.4214 on 3008 degrees of freedom Multiple R-squared: 0.09874, Adjusted R-squared: 0.09844 Wald-statistic: 320.5 on 1 and Inf DF, p-value: < 2.2e-16

Return to education are estimated at 5.2%



#### **Regression with controls**

Dependent varaible: lwage

#### Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	4.6252589	0.0833168	55.514	< 2e-16
educ	0.0733229	0.0045343	16.171	< 2e-16
black	-0.1614265	0.0244866	-6.592	5.39e-11
south	-0.1104188	0.0177876	-6.208	6.40e-10
smsa	0.1621471	0.0180048	9.006	< 2e-16
exper	0.0885766	0.0080561	10.995	< 2e-16
expersq	-0.0023789	0.0003999	-5.949	3.13e-09
fatheduc	-0.0007076	0.0031278	-0.226	0.821
motheduc	0.0078193	0.0037065	2.110	0.035

Heteroskadasticity robust standard errors used

Residual standard error: 0.3773 on 2211 degrees of freedom
 (790 observations deleted due to missingness)
Multiple R-squared: 0.2663, Adjusted R-squared: 0.2636
Wald-statistic: 836.8 on 8 and Inf DF, p-value: < 2.2e-16</pre>

### Return to education $\approx 7.3\%$





> summary\_rob(lm(lwage~educ+black+south+smsa+exper+expersq+fatheduc+motheduc+IQ+KWW,data=cd))
Dependent varaible: lwage

#### Coefficients:

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	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	4.4909962	0.1266845	35.450	< 2e-16
educ	0.0628479	0.0066013	9.521	< 2e-16
black	-0.0963360	0.0364614	-2.642	0.008319
south	-0.0784613	0.0209866	-3.739	0.000192
smsa	0.1473359	0.0213596	6.898	7.59e-12
exper	0.0903101	0.0110791	8.151	7.21e-16
expersq	-0.0026815	0.0005713	-4.693	2.92e-06
fatheduc	-0.0046165	0.0037345	-1.236	0.216571
motheduc	0.0079514	0.0044815	1.774	0.076211
IQ	0.0018442	0.0008612	2.142	0.032384
KWW	0.0039044	0.0016718	2.335	0.019648

Heteroskadasticity robust standard errors used

Residual standard error: 0.3723 on 1593 degrees of freedom

(1406 observations deleted due to missingness)
Multiple R-squared: 0.2162, Adjusted R-squared: 0.2113
Wald-statistic: 472.1 on 10 and Inf DF, p-value: < 2.2e-16</pre>

Return to education  $\approx 6.28\%$ 



#### Retrun to education

# Returns to education are approximately 6% Confidence interval is [0.050, 0.076]



- nearc2 and nearc4 are indicator denoting whether the individual lives close to a 2 years and a 4 years college
- □ Are nearc2 and nearc4 valid instruments?
  - Relevance
  - Exogeneity



#### Relevance – First Stage Regressions

> summary\_rob(lm(educ~nearc2+nearc4+black+south+smsa+exper+expersq+fatheduc+motheduc+IQ+KWW, data=cd))
Dependent varaible: educ

#### Coefficients:

\_\_\_

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	9.886683	0.429152	23.038	< 2e-16
nearc2	-0.007681	0.077462	-0.099	0.921020
nearc4	0.231521	0.087863	2.635	0.008495
black	0.681182	0.140342	4.854	1.33e-06
south	0.129108	0.081699	1.580	0.114239
smsa	0.005987	0.094456	0.063	0.949470
exper	-0.553602	0.037193	-14.885	< 2e-16
expersq	0.012706	0.001953	6.504	1.04e-10
fatheduc	0.053552	0.014959	3.580	0.000354
motheduc	0.049008	0.017066	2.872	0.004137
IQ	0.022029	0.003291	6.694	3.01e-11
KWW	0.112311	0.005781	19.427	< 2e-16

Heteroskadasticity robust standard errors used

Residual standard error: 1.48 on 1592 degrees of freedom
 (1406 observations deleted due to missingness)
Multiple R-squared: 0.5762, Adjusted R-squared: 0.5733
Wald-statistic: 3277 on 11 and Inf DF, p-value: < 2.2e-16</pre>

```
wtest(lm1, testcoef=c("nearc2", "nearc4"))
Wald test
```

```
Null hypothesis:
nearc2 = 0
```

nearc4 = 0

q W pvalue 2 7.037288 **0.0296396** 

> The null hypothesis that nearc2 and nearc4 do **not enter** the first stage regression is rejected at 5%



#### TSLS

coeftest(ivreg(lwage~educ+black+south+smsa+exper+expersq+fatheduc+motheduc+IQ+KWW|

.+nearc2+nearc4-educ , data=cd), vcov.=vcovHC)

Estimate Std. Error t value Pr(>|t|) (Intercept) 4.6448384 0.9213975 5.0411 5.157e-07 \*\*\* educ 0.0474293 0.0916678 0.5174 0.6049458 black -0.0857098 0.0741675 -1.1556 0.2480081 -0.0770074 0.0227986 -3.3777 0.0007485 \*\*\* south 0.1484669 0.0223741 6.6357 4.415e-11 \*\*\* smsa 0.0817947 0.0519974 1.5731 0.1159051 exper -0.0024872 0.0013017 -1.9107 0.0562248 . expersq fatheduc -0.0037760 0.0062892 -0.6004 0.5483230 motheduc 0.0086968 0.0063199 1.3761 0.1689832 0.0021872 0.0021802 1.0032 0.3158995 IQ 0.0056362 0.0104880 0.5374 0.5910661 KWW \_ \_ \_

